

Invariant Histograms

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Abstract

We introduce and study a Euclidean-invariant distance histogram function for curves. For sufficiently regular plane curves, we prove convergence of the cumulative distance histograms based on discretizing the curve by either uniformly spaced or randomly chosen sample points. Robustness of the curve histogram function under noise and pixelization of the curve is also established. We argue that the histogram function serves as a simple, noise-resistant shape classifier for regular curves under the Euclidean group of rigid motions. Extensions of the underlying ideas to higher dimensional submanifolds, as well as to area histogram functions invariant under the group of area-preserving affine transformations are presented.

1 Introduction.

Given a finite set of points contained in \mathbb{R}^n , equipped with the usual Euclidean metric, consider the histogram formed by the mutual distances between all distinct pairs of points. An interesting question, first studied in depth by Boutin and Kemper, [5, 6], is to what extent the distance histogram uniquely determines the point set. Clearly, if the point set is subjected to a rigid motion — a combination of translations, rotations and reflections — the interpoint distances will not change, and so two finite, rigidly equivalent sets have identical distance histograms. However, there do exist finite point sets that have identical histograms but are not rigidly equivalent. (The reader new to the subject may enjoy trying to find an example before proceeding further.) Nevertheless, Boutin and Kemper proved that, in a wide range of situations, the set of such counterexamples is “small” — more precisely, it forms an algebraic variety of lower dimension in the space of all point configurations. Thus, one can say that, usually, the distance histogram uniquely determines a finite point set up to rigid equivalence. This motivates the use of the distance histogram as a simple, robust, noise resistant signature that can be used to distinguish most rigidly inequivalent finite point sets, particularly those that arise as landmark points on an object in a digital image.

The goal of this paper is to develop a comparable distance histogram function for continua — specifically curves, surfaces, and even higher dimensional submanifolds of Euclidean spaces. Most of the paper, including all proofs, will concentrate on the simplest scenario:

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a “fully regular”, as defined below, bounded plane curve. (Regularity does allow corners, and so, in particular, includes polygons.) We will approach this problem using the following strategy. We first sample the curve using a finite number of points, and then compute the distance histogram of the sampled point set. We are able to prove, that, in the limit as the curve becomes densely sampled, the appropriately scaled cumulative distance histograms converge to an explicit function that we name the *global curve histogram function*. Alternatively, computing the histogram of distances from the sample points to a fixed point on the curve leads, in the limit, to a *local curve histogram function*, from which the global version can be obtained by integration with respect to arc length. Convergence of both the local and global sample histograms is rigorously established under three scenarios: first, for uniformly sampled points that are separated by a common arc length distance; second, for randomly sampled points; and, finally, for “pixelized curves”, where we discretize using pixel coordinates in a digital representation of the curve.

The global curve histogram function can be computed directly through an explicit arc length integral. By construction, it is invariant under rigid motions. Hence, a basic question arises: does the histogram function uniquely determine the curve up to rigid motion? There is ample evidence that, under appropriate hypotheses, such a result is true. However, we have been unable to find a complete proof, and so must state this as an open conjecture. If true, the global curve histogram function, as approximated by its sampled point histograms, can be unambiguously employed as an elementary, robust classifier for distinguishing shapes in digital images. As such, it can serve as a simpler substitute for the joint invariant signatures proposed in [17]. Extensions of these ideas to subsets of higher dimensional Euclidean spaces, or even general metric spaces, are immediate. Proofs of convergence in sufficiently regular situations can be established along the same lines as the planar curve case treated here. An interesting question is whether the distance histogram can be used to distinguish objects of differing dimensions.

Following Boutin and Kemper, [5], we also consider area histograms formed by triangles whose corners lie in a finite point set. Area histograms are invariant under the group of area-preserving affine (or equi-affine) transformations. We exhibit a limiting area histogram function for plane curves that is also equi-affine invariant, and propose a similar conjecture. Generalizations to other transformation groups, e.g., similarity, projective, conformal, etc., of interest in image processing and elsewhere, [11, 18], are worth pursuing. The corresponding discrete histograms will be based on suitable joint invariants — for example, area and volume cross ratios in the projective case — which can be systematically classified by the equivariant method of moving frames, [17]. The corresponding limiting curve and submanifold histograms can then be constructed by an analogous procedure; details will be developed elsewhere.

Our study of distance and other histograms has been motivated in large part by the potential applications to object recognition, shape classification, and geometric modeling. Discrete histograms appear in a broad range of powerful image processing algorithms, for shape representation and classification, [1, 24], for image enhancement, [24, 25], in the scale-invariant feature transform (SIFT) [12, 20], in object-based query methods, [23], and as integral invariants, [13, 21]. Local distance histograms underly the method of shape contexts, [3], while their global counterparts are utilized in the method of shape distributions, [19]. Distance histograms provide lower bounds for Gromov–Hausdorff and Gromov–Wasserstein distances for shape matching and comparison, [14, 15].

2 Distance Histograms.

Let us first review the results of Boutin and Kemper, [5, 6], on distance histograms defined by finite point sets. For this purpose, our initial context is a general metric space V , equipped with distance function $d(z, w) \geq 0$, for $z, w \in V$, satisfying the usual axioms.

Definition 1. The *distance histogram* of a finite set of points $P = \{z_1, \dots, z_n\} \subset V$ is the function $\eta = \eta_P: \mathbb{R}^+ \rightarrow \mathbb{N}$ defined by

$$\eta(r) = \#\{(i, j) \mid 1 \leq i < j \leq n, d(z_i, z_j) = r\} \quad \text{for } r \geq 0. \quad (2.1)$$

In this paper, we will restrict our attention to the simplest situation, when $V = \mathbb{R}^m$ is endowed with the usual Euclidean metric, so $d(z, w) = \|z - w\|$. We say that two subsets $P, Q \subset V$ are *rigidly equivalent*, written $P \simeq Q$, if we can obtain Q by applying an isometry to P . In Euclidean geometry, isometries are *rigid motions* — the translations, rotations, and reflections, generating the Euclidean group, [27]. Clearly, any two rigidly equivalent finite subsets have identical distance histograms. Boutin and Kemper’s main result is that the converse is, in general, false, but is true for a broad range of generic point configurations.

Theorem 2. Let $\mathcal{P}^{(n)} = \mathcal{P}^{(n)}(\mathbb{R}^m)$ denote the space of finite (unordered) subsets $P \subset \mathbb{R}^m$ of cardinality $\#P = n$. If $n \leq 3$ or $n \geq m + 2$, then there is a Zariski dense open subset $\mathcal{R}^{(n)} \subset \mathcal{P}^{(n)}$ with the property that two subsets $P, Q \in \mathcal{R}^{(n)}$ have identical distance histograms, $\eta_P = \eta_Q$, if and only if they are rigidly equivalent: $P \simeq Q$.

In other words, for the indicated ranges of n , unless the points are constrained by a certain algebraic equation, and so are “non-generic”, the distance histogram uniquely determines the point configuration up to a rigid motion. Interestingly, the simplest counterexample is not provided by the corners of a regular polygon. For example, the corners of a unit square have 4 side distances of 1 and 2 diagonal distances of $\sqrt{2}$, i.e., the distance histogram has values $\eta(1) = 4$, $\eta(\sqrt{2}) = 2$, while $\eta(r) = 0$ for $r \neq 1, \sqrt{2}$. Moreover, this is the only possible way to arrange four points with the given distance histogram. A simple non-generic configuration is provided by the corners of the kite and trapezoid quadrilaterals shown in Figure 1. Although clearly not rigidly equivalent, both point configurations have the same distance histogram, with nonzero values $\eta(\sqrt{2}) = 2$, $\eta(2) = 1$, $\eta(\sqrt{10}) = 2$, $\eta(4) = 1$.

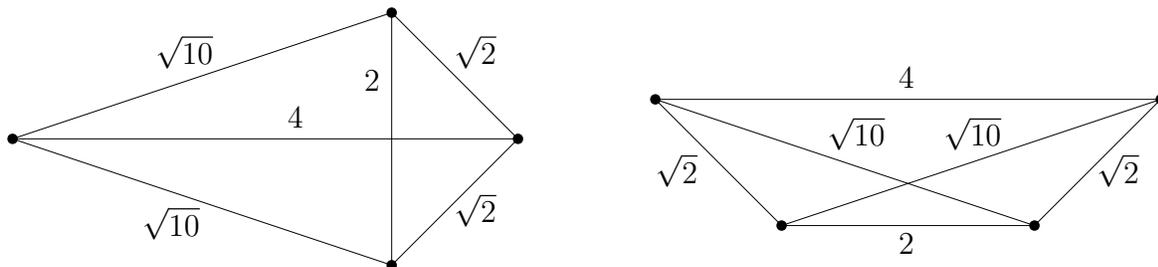


Figure 1: Kite and Trapezoid.

An intriguing one-dimensional counterexample, discovered in [4], is provided the two sets $P = \{0, 1, 4, 10, 12, 17\} \subset \mathbb{R}$ and $Q = \{0, 1, 8, 11, 13, 17\} \subset \mathbb{R}$, which have identical distance histograms, but are clearly not rigidly equivalent.

For our analysis, it will be more convenient to work with the (renormalized) *cumulative distance histogram function* (also known as the *histogram distribution function*)

$$\Lambda_P(r) = \frac{1}{n} + \frac{2}{n^2} \sum_{s \leq r} \eta_P(s) = \frac{1}{n^2} \# \{ (i, j) \mid d(z_i, z_j) \leq r \}. \quad (2.2)$$

We note that we can recover the usual distance histogram (2.1) via

$$\eta(r) = \Lambda_P(r) - \Lambda_P(r - \delta) \quad \text{for sufficiently small } \delta \ll 1. \quad (2.3)$$

We further introduce a *local distance histogram function* that counts the fraction of points in P that are within a specified distance r of a given point $z \in \mathbb{R}^m$ (not necessarily on C):

$$\lambda_P(r, z) = \frac{1}{n} \# \{ z_j \in P \mid d(z, z_j) \leq r \} = \frac{1}{n} \#(P \cap B_r(z)), \quad (2.4)$$

where

$$B_r(z) = \{ v \in V \mid d(v, z) \leq r \}, \quad S_r(z) = \partial B_r(z) = \{ v \in V \mid d(v, z) = r \}, \quad (2.5)$$

denote, respectively, the ball of radius r (in the plane, the disk) centered at the point z and its bounding sphere (circle). Observe that we recover the cumulative histogram function (2.2) by averaging its localization

$$\Lambda_P(r) = \frac{1}{n} \sum_{z \in P} \lambda_P(r, z) = \frac{1}{n^2} \sum_{z \in P} \#(P \cap B_r(z)). \quad (2.6)$$

In this paper, we are primarily interested in the case when the points lie on a curve. Until the final section, we restrict our attention to plane curves: $C \subset V = \mathbb{R}^2$. A finite subset $P \subset C$ will be called a set of *sample points* on the curve. We will assume throughout that the curve C is bounded, rectifiable, and, usually (although not necessarily), closed. Further mild regularity conditions will be introduced below. We use $z(s)$ to denote the arc length parametrization of C , measured from some base point $z(0) \in C$. Let

$$L = l(C) = \int_C ds < \infty \quad (2.7)$$

denote the curve's length, which we always assume to be finite.

Our aim is to study the limiting behavior of the cumulative histogram functions constructed from more and more densely chosen sample points. It turns out that, under reasonable assumptions, the discrete histogram functions do converge, and the limiting function can be characterized as follows.

Definition 3. Given a curve $C \subset V$, the *local distance histogram function* based at a point $z \in V$ is

$$h_C(r, z) = l(C \cap B_r(z)), \quad (2.8)$$

i.e., the total length of those parts of the curve contained within the disk of radius r centered at z . The *global distance histogram function* of C is obtained by averaging the local version over the curve:

$$H_C(r) = \frac{1}{L} \int_C h_C(r, z(s)) ds. \quad (2.9)$$

Observe that formula (2.9) is invariant under rigid motions, and hence two curves that are rigidly equivalent have identical global histogram functions. An interesting question, which we consider in some detail towards the end of the paper, is whether the global histogram function uniquely characterizes the curve up to rigid equivalence.

Modulo the definition of “fully regular”, to be presented in the following section, our main result can be stated as follows.

Theorem 4. *Let C be a fully regular plane curve of length L . Then, for both uniformly and randomly chosen sample points $P \subset C$, the discrete local and global histogram functions converge to their continuous counterparts:*

$$\lambda_P(r, z) \longrightarrow \frac{1}{L} h_C(r, z), \quad \Lambda_P(r) \longrightarrow \frac{1}{L} H_C(r), \quad (2.10)$$

as the number of sample points goes to infinity.

3 Local Histogram Functions.

Our proof of Theorem 4 begins by establishing the convergence of local histogram functions. In this section, we work under the assumption that the sample points are uniformly spaced along the curve.

Let us recall some basic terminology concerning plane curves, mostly taken from Guggenheimer’s book [10]. We will assume throughout that $C \subset \mathbb{R}^2$ has a piecewise C^2 arc length parametrization $z(s)$, where $0 \leq s \leq L$ belongs to a bounded closed interval, with $L = l(C) < \infty$ being the overall length. The curve is always assumed to be *simple*, meaning that there are no self-intersections, and either *closed*, i.e., a Jordan curve, or what we will call a *curve segment* that has distinct endpoints $z(0) \neq z(L)$. By convention, we will also designate a single point to be a segment of length 0. We use $t(s) = z'(s)$ to denote the unit tangent, and¹ $\kappa(s) = z'(s) \wedge z''(s)$ to denote the signed curvature at the point $z(s)$. Under our assumptions, both $t(s)$ and $\kappa(s)$ have left and right hand limiting values at their finitely many discontinuities. A point $z(s) \in C$ where either the tangent or curvature is not continuous will be referred to as a *corner*.

A closed curve is called *convex* if it bounds a convex region in the plane. A curve segment is *convex* if the region bounded by it and the straight line segment connecting its endpoints is a convex region. A curve segment is called a *spiral arc* if the curvature function $\kappa(s)$ is continuous, monotone (but not necessarily strictly monotone), and of one sign, i.e., either $\kappa(s) > 0$ or $\kappa(s) < 0$ except we allow the curvature to vanish at one of the endpoints. By convention, straight lines will also be considered to be spiral arcs.

¹The symbol \wedge is the two-dimensional cross product: $v \wedge w = v_1 w_2 - v_2 w_1$.

Definition 5. A plane curve is called *regular* if it is piecewise C^2 and the union of a finite number of convex spiral arcs.

Thus, any regular curve has only finitely many *corners*, finitely many *inflection points*, where the curvature has an isolated zero, and finitely many *vertices*, meaning points where the curvature has a local maximum or minimum, but is not locally constant (i.e., not on a line segment or circular arc). In particular, polygons are regular, as are piecewise circular curves, also known as *biarcs*, [16]. However, our terminological convention is that polygons and biarcs have corners, not vertices! Examples of irregular curves include the graph of the infinitely oscillating function $y = x^5 \sin 1/x$ near $x = 0$, and the non-convex spiral arc $r = e^{-\theta}$ for $0 \leq \theta < \infty$, expressed in polar coordinates.

The reason for restricting our attention to regular curves is the following result.

Theorem 6. *If C is a regular plane curve, then there is a positive integer $0 < m_C < \infty$ such that the curve's intersection, $C \cap B_r(z)$, with any disk having center $z \in C$ and radius $r > 0$ consists of at most m_C connected segments. The minimal value of m_C will be called the circular index of C .*

Proof: This is an immediate consequence of a Theorem of Vogt, [26] — see also Exercise 3-3.11 on page 53 of [10] — that states that a convex spiral arc and a circle intersect in at most 3 points. Thus, $m_C \leq 3m$, where $m < \infty$ is the number of convex spiral arcs (including straight lines) needed to form C . *Q.E.D.*

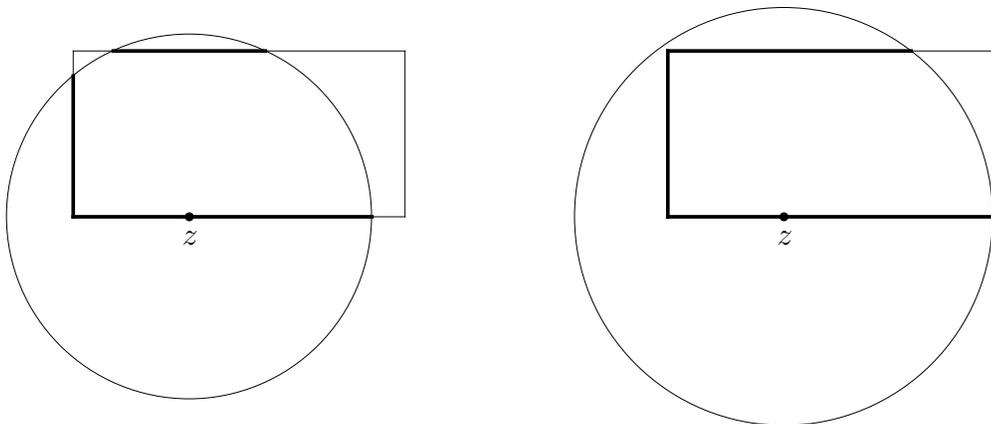


Figure 2: Intersections of a Rectangle and a Disk.

Example 7. Let C be a rectangle. A disk $B_r(z)$ centered at a point $z \in C$ will intersect the rectangle in either one or two connected segments; see Figure 2. Thus, the circular index of a rectangle is $m_C = 2$.

For each positive integer $n > 0$, let $P_n \subset C$ denote a collection of uniform sample points separated by a common arc length spacing $\Delta l = L/n$. Observe that when C is a closed curve, n equals the number of sample points, while when C is a non-closed curve segment, there are $n + 1$ sample points including the two endpoints,.

Lemma 8. *Let C be a regular curve. Then, for each $z \in C$ and $r > 0$, the corresponding local histogram functions based on uniformly spaced sample points $P_n \subset C$ converge:*

$$\lambda_n(r, z) = \lambda_{P_n}(r, z) \longrightarrow \frac{1}{L} h_C(r, z) \quad \text{as } n \rightarrow \infty. \quad (3.1)$$

Proof. We will prove convergence by establishing the bound

$$|h_C(r, z) - L \lambda_n(r, z)| \leq m_C \Delta l, \quad (3.2)$$

where m_C is the circular index of C .

By assumption, since $z \in C$, the intersection $C \cap B_r(z) = S_1 \cup \dots \cup S_k$ consists of $1 \leq k \leq m_C$ connected segments whose endpoints lie on the bounding circle $S_r(z)$. Since the sample points are uniformly spaced by $\Delta l = L/n$, the number of sample points n_i contained in an individual segment S_i can be bounded by

$$(n_i - 1) \Delta l \leq l(S_i) < (n_i + 1) \Delta l.$$

Summing over all segments, and noting that

$$\sum_i n_i = \#(P_n \cap B_r(z)) = n \lambda_n(r, z), \quad \sum_i l(S_i) = l(C \cap B_r(z)) = h_C(r, z),$$

we deduce that

$$L \lambda_n(r, z) - k \Delta l \leq h_C(r, z) < L \lambda_n(r, z) + k \Delta l,$$

from which (3.2) follows. *Q.E.D.*

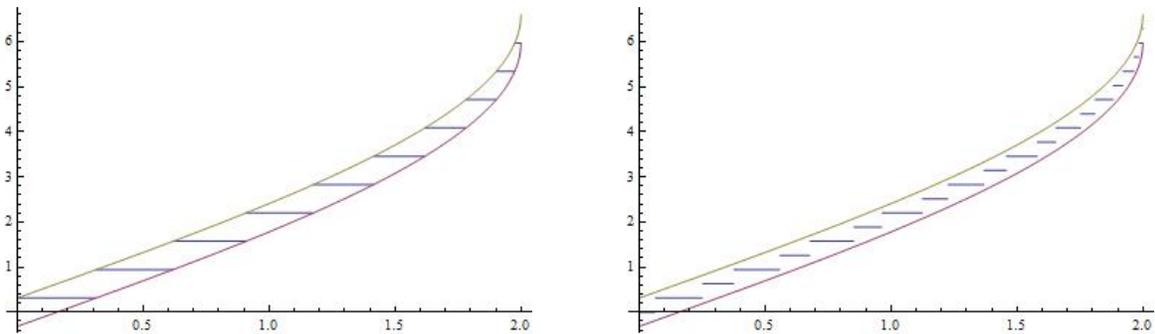


Figure 3: Local Histogram Functions for a Circle

Example 9. Let C be a circle of radius 1. A set of n evenly spaced sample points $P_n \subset C$ form the corners of a regular n -gon. Using the identification $\mathbb{R}^2 \simeq \mathbb{C}$, the polygon's cumulative histogram function is given by

$$\lambda_n(r, z) = \frac{1}{n} \# \{ 1 \leq j \leq n \mid |e^{2\pi i j/n} - z| < r \}.$$

On the other hand, the local histogram function (2.8) for a circle is easily found to have the explicit form

$$h_C(r, z) = 2 \cos^{-1} \left(1 - \frac{1}{2} r^2 \right), \quad (3.3)$$

which, by symmetry, is independent of the point $z \in C$.

In Figure 3, we plot the discrete cumulative histogram $L \lambda_n(r, z) = 2 \pi \lambda_n(r, z)$ for $n = 20$, along with the bounds $h_C(r, z) \pm \Delta l$ coming from (3.2), which reflects the fact that a circle has circular index $m_C = 1$. In the first plot, the center z coincides with a data point, while the second takes z to be a distance .01 away, as measured along the circle. Observe that the discrete histogram stays within the bounds at all radii, in accordance with our result.

4 Global Histogram Functions.

We now turn our attention to the convergence of the global histogram function. Again, we work under the preceding regularity assumptions, and continue to focus our attention on the case of uniformly sampled points $P_n \subset C$.

The local histogram function $h(s) = h_C(r, z(s))$ is clearly continuous as a function of s . Thus, we can approximate the global histogram integral (2.9) as a Riemann sum based on the evenly spaced data points:

$$H_C(r) = \frac{1}{L} \int_C h_C(r, z(s)) ds \approx \frac{1}{L} \sum_{z \in P_n} h_C(r, z) \Delta l. \quad (4.1)$$

Since C has finite length, $\Delta l = L/n \rightarrow 0$ as $n \rightarrow \infty$, and so the Riemann sums converge.

On the other hand, (3.1) implies that the local histogram function can be approximated by the (rescaled) cumulative point histogram $L \lambda_n(r, z)$, and hence we should be able to approximate the Riemann sum in turn by

$$\frac{1}{L} \sum_{z \in P_n} L \lambda_n(r, z) \Delta l = \frac{L}{n} \sum_{z \in P_n} \lambda_n(r, z) = L \Lambda_n(r), \quad (4.2)$$

using the first equality of (2.6). This will imply the global convergence result (2.9).

To bound the approximation error, we proceed as follows. Recall, [2], that the *total variation* of a function $f(x)$ on the interval $[a, b]$ is defined as

$$\mathcal{V}_a^b[f] = \sup \sum_{i=1}^{n-1} |f(x_{i+1}) - f(x_i)|, \quad (4.3)$$

where the supremum is taken over all possible subdivisions $a \leq x_1 < x_2 < \dots < x_n \leq b$. If $f(x)$ is monotone, then $\mathcal{V}_a^b[f] = |f(b) - f(a)|$. If $f(x)$ is piecewise continuously differentiable, then $\mathcal{V}_a^b[f] = \int_a^b |f'(x)| dx$ is the integral of the absolute value of its derivative. A function is said to have *bounded variation* if $\mathcal{V}_a^b[f] < \infty$. In particular, all piecewise C^1 functions are of bounded variation, as are all functions that are monotone on a finite number of subintervals. Indeed, any function of bounded variation is the difference of two monotone functions. The *cumulative property* states that if $a \leq b \leq c$, then $\mathcal{V}_a^c[f] = \mathcal{V}_a^b[f] + \mathcal{V}_b^c[f]$.

Curiously, we were unable to find an exact statement and proof in the literature of the following simple result concerning the numerical approximation of the integral of a function of bounded variation.

Theorem 10. *If $f(x)$ is of bounded variation on $[a, b]$, then the difference between the integral $\int_a^b f(x) dx$ and any approximating Riemann sum is bounded in absolute value by $\mathcal{V}_a^b[f] \Delta x$, where Δx is the maximal subinterval length appearing in the Riemann sum.*

Proof: Let $a \leq x_1 < x_2 < \dots < x_n \leq b$ be a partition, and

$$R = \sum_{i=1}^{n-1} f(y_i) \Delta x_i$$

a corresponding Riemann sum, where $x_i \leq y_i \leq x_{i+1}$ and $\Delta x_i = x_{i+1} - x_i$. On the i -th subinterval $[x_i, x_{i+1}]$, let x_i^+ and x_i^- be the points at which $f(x)$ achieves, respectively, its maximum and minimum values. Then

$$f(x_i^-) \Delta x_i \leq \int_{x_i}^{x_{i+1}} f(x) dx \leq f(x_i^+) \Delta x_i,$$

and hence

$$\left| \int_{x_i}^{x_{i+1}} f(x) dx - f(y_i) \Delta x_i \right| \leq |f(x_i^*) - f(y_i)| \Delta x_i \leq \mathcal{V}_{x_i}^{x_{i+1}}[f] \Delta x_i,$$

where $x_i^* = x_i^+$ or x_i^- is the point that maximizes $|f(x) - f(y_i)|$ on the interval. Therefore, by the cumulative property of the total variation,

$$\left| \int_a^b f(x) dx - \sum_{i=1}^{n-1} f(y_i) \Delta x_i \right| \leq \sum_{i=1}^{n-1} \mathcal{V}_{x_i}^{x_{i+1}}[f] \Delta x_i \leq \mathcal{V}_a^b[f] \Delta x,$$

completing the proof. *Q.E.D.*

Recall that, given a C^1 curve C , the *parallel curve* $\Pi_r(C)$ of distance $r > 0$ is the locus of points that are a distance r away from C , as measured along the normal direction. Since there are two normal directions at each point, $\Pi_r(C)$ consists of two connected, not necessarily simple, curves. If C is only piecewise C^1 , then $\Pi_r(C)$ is the union of the r parallel curves to each C^1 segment combined with circular arcs of radius r centered at each corner that connect the endpoints of the two parallel curve segments.

Definition 11. A regular curve C is called *fully regular* if, for each $r > 0$, the intersection $C \cap \Pi_r(C)$ has finitely many connected components.

Not all regular curves are fully regular. For example, it is possible to slightly deform part of a circle of radius $\frac{1}{2}r$ to produce a smooth convex curve that intersects its parallel curve of distance r infinitely often. On the other hand, most regular curves, including all polygons and biarcs, are fully regular.

Theorem 12. *If C is a fully regular curve of length L , then the function $h(s) = h_C(r, z(s))$ is continuous and of bounded variation on $[0, L]$.*

This result is a consequence of the following formula for the derivative of the local histogram function.

Proposition 13. *Let $z = z(s) \in C$. Let $y_1 = z(s_1), \dots, y_k = z(s_k)$ denote the points of intersection of C with the circle $S_r(z)$ of radius r centered at z . For each $j = 1, \dots, k$, let l_j denote the line through z and y_j . Let θ_j denote the angle between l_j and the tangent vector $t(s) = z'(s)$ to C at z in the direction of increasing s . Let φ_j denote the angle between the line l_j and the tangent vector (not necessarily in the direction of increasing s) to C at y_j that points outside the circle. Then*

$$\frac{d}{ds} h(r, z(s)) = \sum_{j=1}^k \frac{\cos \theta_j}{\cos \varphi_j}. \quad (4.4)$$

Proof: Referring to Figure 4, note first that $\cos \varphi_j = 0$ if and only if the curve C is tangent to the circle $S_r(z)$ at the intersection point y_j . In this case, the line l_j is normal to C at y_j , and hence $z \in \Pi_r(C)$, the parallel curve at distance r . We exclude these configurations from the remainder of the argument.

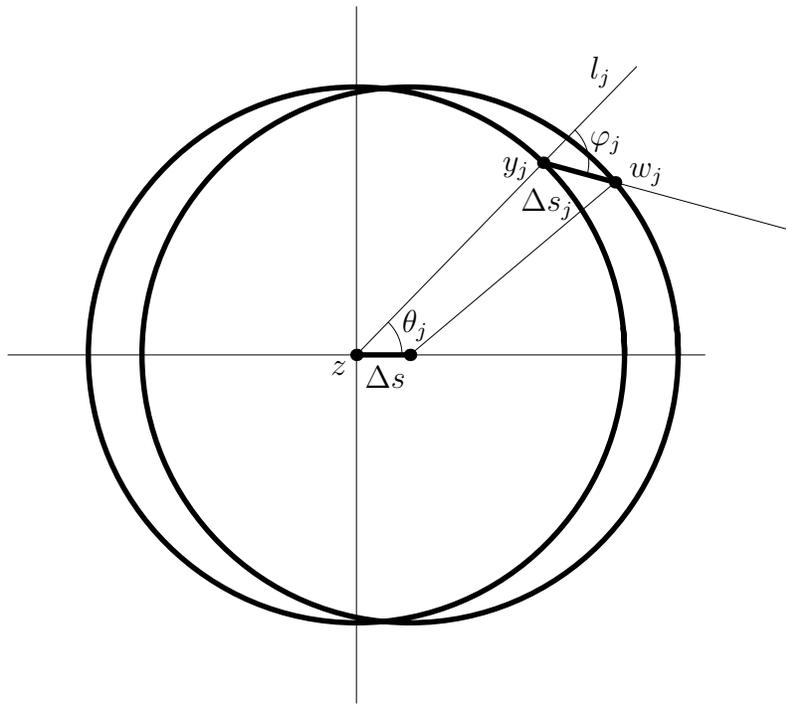


Figure 4: Calculating the Derivative of the Distance Function.

The calculation is simplified by identifying $\mathbb{R}^2 \simeq \mathbb{C}$, so $z = x + iy$. We place $z = z(s) = 0$ at the origin, with the tangent vector to C at z in the direction of the positive x axis. Let $0 < \Delta s \ll 1$. To first order, $z(s + \Delta s) = \Delta s + \dots$. Moreover, as Δs varies, the intersection point $y_j = z(s_j) = r_j e^{i\theta_j}$ moves to a point

$$w_j = z(s_j + \Delta s_j) = y_j + e^{i(\theta_j - \varphi_j)} \Delta s_j + \dots = e^{i\theta_j} (r_j + e^{-i\varphi_j} \Delta s_j) + \dots$$

on the circle $S_r(z(s + \Delta s))$, satisfying the algebraic equation

$$\begin{aligned}
r^2 &= |w_j - z(s + \Delta s)|^2 = |e^{i\theta_j}(r_j + e^{-i\varphi_j}\Delta s_j) - \Delta s|^2 + \dots \\
&= |r_j + e^{-i\varphi_j}\Delta s_j - e^{-i\theta_j}\Delta s_j|^2 + \dots \\
&= (r_j + \Delta s_j \cos \varphi_j - \Delta s \cos \theta_j)^2 + (-\Delta s_j \sin \varphi_j - \Delta s \sin \theta_j)^2 + \dots \\
&= r^2 - 2r\Delta s \cos \theta_j + 2r\Delta s_j \cos \varphi_j + \dots .
\end{aligned}$$

Solving, to leading order,

$$\frac{\Delta s_j}{\Delta s} = \frac{\cos \theta_j}{\cos \varphi_j} + \dots .$$

Taking the limit as $\Delta s \rightarrow 0$ and then summing over the points y_j in the intersection completes the proof. *Q.E.D.*

Proof of Theorem 12: As a consequence of (4.4), the derivative $h'(s)$ is defined and nonzero, unless either $z = z(s)$ is a corner of C , or $\cos \varphi_j = 0$, or $\cos \theta_j = 0$. As noted above, the second possibility means that the point $z \in \Pi_r(C)$. The third possibility means that the point $y_j \in C \cap S_r(z)$ lies a distance r away from $z \in C$, as measured along the normal direction at z , and hence $y_j \in \Pi_r(C)$. Thus, by our fully regular hypothesis, $h'(s)$ is defined and nonzero except on a finite number of points and/or connected subintervals. We conclude that $h(s)$ is continuous, piecewise monotone, and hence of bounded variation. *Q.E.D.*

Let $V = \mathcal{V}_0^L[h]$ denote the total variation of $h(s) = h_C(r, z(s))$. Theorem 10 implies that the error in the Riemann sum approximation (4.1) can be bounded by

$$\left| \frac{1}{n} \sum_{z \in P_n} h_C(r, z) - \frac{1}{L} \int_C h_C(r, z(s)) ds \right| \leq V \Delta l.$$

On the other hand, (3.2) and the triangle inequality imply that

$$\left| \frac{1}{n} \sum_{z \in P_n} L \lambda_n(r, z) - \frac{1}{n} \sum_{z \in P_n} h_C(r, z) \right| \leq m_C \Delta l. \quad (4.5)$$

Combining these two estimates, we find that the difference between the continuous and discrete global histogram functions can be bounded by

$$|L \Lambda_n(r) - H_C(r)| = \left| \frac{1}{n} \sum_{z \in P_n} L \lambda_n(r, z) - \frac{1}{L} \int_C h_C(r, z(s)) ds \right| \leq (m_C + V) \Delta l. \quad (4.6)$$

Thus, under our hypotheses, the convergence to the global histogram function is first order in the interpoint spacing Δl .

Example 14. Let C be a unit square. Measuring the arc length s along the square starting at a corner, the local histogram function $h_r(s) = h_C(r, z(s))$ can be explicitly constructed

using elementary geometry, distinguishing several different configurations. For $0 \leq s \leq \frac{1}{2}$:

$$h_r(s) = \begin{cases} 2r, & 0 \leq r \leq s, \\ s + r + \sqrt{r^2 - s^2}, & s \leq r \leq 1 - s, \\ 1 + \sqrt{r^2 - s^2} + \sqrt{r^2 - (1 - s)^2}, & 1 - s \leq r \leq 1, \\ 1 + 2\sqrt{r^2 - 1} + \sqrt{r^2 - s^2} + \sqrt{r^2 - (1 - s)^2}, & 1 \leq r \leq \sqrt{1 + s^2}, \\ s + 2 + \sqrt{r^2 - 1} + \sqrt{r^2 - (1 - s)^2}, & \sqrt{1 + s^2} \leq r \leq \sqrt{1 + (1 - s)^2}, \\ 4 & \sqrt{1 + (1 - s)^2} \leq r, \end{cases} \quad (4.7)$$

while other values follow from the fact that $h_r(s)$ is both 1 periodic and even:

$$h_r(1 - s) = h_r(s) = h_r(1 + s).$$

Integration with respect to arc length produces the global histogram function

$$H_C(r) = \begin{cases} 2r + \left(\frac{1}{2}\pi - 1\right)r^2, & r < 1, \\ 2 - r^2 + 4\sqrt{r^2 - 1} + r^2 \left(\sin^{-1} \frac{1}{r} - \cos^{-1} \frac{1}{r}\right), & 1 \leq r < \sqrt{2}, \\ 4 & r \geq \sqrt{2}. \end{cases} \quad (4.8)$$

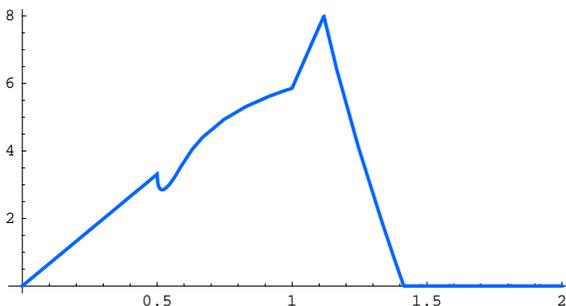


Figure 5: Total Variation of the Local Histogram Function of a Unit Square.

The total variation of $h_r(s)$ is obtained by integrating the absolute value of its derivative. In Figure 5, we plot the variation as a function of r . (The little dip after $r = .5$ is genuine, and the reader may enjoy seeking a geometrical explanation.) The maximal variation, $V_{max} = 8$, occurs at $r = \sqrt{5}/2$, where $h_{\sqrt{5}/2}(s)$ varies monotonically from 3 to 4 and then back four times as we go around the square. On the other hand, according to Example 7, $m_C = 2$. Thus, an overall bound (4.6) for the error in $L \lambda_P(r, z)$, valid for all r , is $(m_C + V_{max}) \Delta l = 10 \Delta l$.

A Mathematica plot of the histogram along with the bounds Δl and $2 \Delta l$ is shown in Figure 6. The first plot gives the local histogram $L \lambda_n(r, z)$; the second plots the global histogram $L \Lambda_n(r)$, both based on $n = 20$ points. Observe that the discrete histogram, in fact, stays within Δl of the curve histogram, a much tighter bound than we derived analytically. Interestingly, a similar bound appears to hold in all the examples we have checked so far.

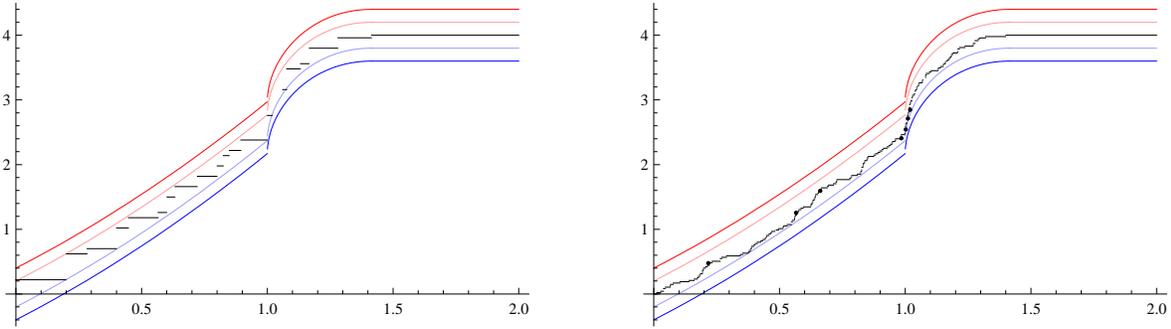


Figure 6: Local and Global Histogram Bounds for a Square

5 Random Point Distributions.

We have thus far proved, under suitable regularity hypotheses, convergence of the discrete histogram functions, both local and global, when constructed from uniformly spaced sample points along the curve. However, in practice, it may be difficult to ensure precise uniform spacing of the sample points. For example, if C is an ellipse, then this would require evaluating n elliptic integrals. Hence, for the purposes of practical shape analysis, we need to examine more general methods of histogram creation. In this section, we analyze the case of randomly chosen sample points $P_n \subset C$, and statistically evaluate the resulting approximation to the global histogram function.

In this case, we view the local discrete histogram function $\lambda_n(r, z)$ as a random variable representing the fraction of the points z_j that lie within a circle of radius r centered at the point z . The expected value of this quantity will represent the probability that a point on the curve C lies inside the disk $B_r(z)$, and so

$$E[\lambda_n(r, z)] = \frac{1}{L} l(C \cup B_r(z)) = \frac{1}{L} h_C(r, z). \quad (5.1)$$

Now, our primary quantity of interest is the global histogram function $H_C(r)$, and so we seek a statistical variable with that as its expectation. Consider the function

$$H_n(r) = L \Lambda_n(r) = \frac{L}{n^2} \sum_i \#(P \cap B_r(z_i)) = \frac{L}{n^2} \sum_i \sum_{j \neq i} \sigma_{i,j}(r). \quad (5.2)$$

Here

$$\sigma_{i,j}(r) = \begin{cases} 1, & d(z_i, z_j) \leq r, \\ 0, & d(z_i, z_j) > r. \end{cases}$$

so that $\sigma_{i,j}(r) = 1$ if and only if the points are within a distance r of each other — precisely when $z_i \in B_r(z_j)$ or, equivalently, $z_j \in B_r(z_i)$. The expectation of this variable will be the probability that, given a starting point, any other point will be within a disk of radius r around it. Since the starting point is random as well, the expected value will be the average of $h_C(r, z_i)$ over all possible values of z_i , which corresponds exactly to $H_C(r)$. We can therefore write

$$H_C(r) = E[H_n(r)] = \frac{L}{n^2} \sum_i \sum_{j \neq i} E[\sigma_{i,j}(r)]. \quad (5.3)$$

With this in hand, we can calculate the variance

$$\begin{aligned}
\text{Var}[H_n(r)] &= E[H_n(r)^2] - E[H_n(r)]^2 \\
&= \frac{L^2}{n^4} \sum_i \sum_{i'} \sum_{j \neq i} \sum_{j' \neq i'} (E[\sigma_{i,j}(r) \sigma_{i',j'}(r)] - E[\sigma_{i,j}(r)] E[\sigma_{i',j'}(r)]) \\
&= \frac{L^2}{n^4} \left(\sum_{i,i',j,j'} E_{i,i',j,j'} - \sum_{i,i',j' \neq i'} E_{i,i',i,j'} - \sum_{i,i',j \neq i} E_{i,i',j,i'} \right),
\end{aligned} \tag{5.4}$$

where

$$E_{i,i',j,j'} = E[\sigma_{i,j}(r) \sigma_{i',j'}(r)] - E[\sigma_{i,j}(r)] E[\sigma_{i',j'}(r)]. \tag{5.5}$$

If z_i, z'_i, z_j, z'_j are independent random variables, then $E_{i,i',j,j'} = 0$, because the two functions inside the first expectation will be independent, and so we can split it into two separate expectations, canceling the second term. Furthermore, since the final two sums in (5.4) are only over three variables (eliminating one variable in the summation by equality), they will be proportional to some constant times n^3 . Thus,

$$\text{Var}[H_n(r)] = O(n^{-1}), \tag{5.6}$$

and hence $H_n(r) = L \Lambda_n(r)$ converges to $H_C(r)$ as $n \rightarrow \infty$, at least in the sense that for any given value of r , the value of $H_n(r)$ will be more likely to be within any given bound as n increases.

Although this probabilistic calculation yields a less stringent error bound in (5.6), it nevertheless provides a practical method for calculating $H_C(r)$ in cases when the extraction of evenly spaced sample points is problematic. However, we still face several important issues. We assumed that P was a set of randomly selected points on the curve, but we did not define how such points should be chosen. Since the calculation of $L \Lambda_n(r)$ essentially reduces to evaluating (2.9), we require the points to be sampled uniformly with respect to arc length. One way to accomplish this in practice is to select points using any convenient parameterization of the curve, but weighting the sample in favor of the points with higher curvature using, for instance, rejection sampling, [8].

6 Pixelization.

One of our motivating goals has been potential applications of distance histograms to the recognition and classification of objects in digitized images. In such situations, it is often more convenient to use the locations of the pixels that the curve passes through to represent the sample points on the shape boundary. This results in two potential errors in the ensuing calculations.

First, the normalization can change, because different pixels will not contain the same length of curve. As long as the *pixel width* w is small, namely $w < 1/\kappa_{max}$, where κ_{max} denotes the maximum curvature on that part of the curve, and the curve intersects the pixel in only one connected segment, then the longest possible curve length is that of a quarter circle, $\frac{1}{2} \pi w$. The shortest possible length depends on the method used to select pixels. In the extreme case, we include all pixels containing an arbitrarily short curve segment — even

if only a single point. So the potential variance in normalization is $\frac{1}{2}\pi w$. However, if the pixelization is very fine then the number of pixels selected to represent the curve will be small compared to the total number of image pixels, and the relative pixel normalizations should not have a noticeable effect.

Second, by using the coordinates of the center of a pixel containing a point, we introduce a possible offset to the location of the sample point of up to a distance

$$\delta = \frac{w}{\sqrt{2}}. \quad (6.1)$$

Thus, for every distance we calculate, we introduce an additional uncertainty of $2\delta = \sqrt{2}w$. If we approximate $h_C(r, z) \approx L\lambda_n(r, z)$, then, even with no statistical error, we still have an uncertainty range of

$$\lambda_n(r - 2\delta, z) < h_C(r, z) < \lambda_n(r + 2\delta, z). \quad (6.2)$$

Since both h and λ_n are monotonically increasing as functions of r , the resulting contribution to the error can thus be bounded by

$$|\lambda_n(r + 2\delta, z) - \lambda_n(r - 2\delta, z)| \approx 4\delta \left| \frac{\partial \lambda_n}{\partial r}(r, z) \right| = 2\sqrt{2}w \left| \frac{\partial \lambda_n}{\partial r}(r, z) \right| \quad \text{for } \delta \ll 1. \quad (6.3)$$

Thus, the error in the pixelized approximation will be proportional to the pixel width. This error will not necessarily go to zero as $n \rightarrow \infty$. However, if we assume that the length of the curve is L finite, then, according to Theorem 17 below, we can bound the number of pixels that the curve passes through. Therefore, if the derivative $\partial h/\partial r$ is bounded in absolute value, then, as $n \rightarrow \infty$, the pixel width $w \rightarrow 0$, and so the local histogram function $\lambda_n(r, z)$ still converges for pixelized images. Similarly, since the global histogram function $H_C(r)$ is just the average of $h_C(r, z)$ along the curve, its pixelized approximation will also converge.

This property is especially important in terms of resistance to noise. For curvature-based techniques of object recognition, as in [7], noise causes large fluctuations in the local curvature-based invariants, which interferes with their use in shape recognition. The distance histogram is largely unaffected by such local noise, and, as we increase the number of points, the effect disappears.

Example 15. Let C be a 2×3 rectangle. In Figure 7, we compare the graphs of the global histogram $H_C(r)/L$ (in black) with the discrete approximations $\Lambda_n(r)$ for $n = 20$ points, using evenly distributed sample points (green), random points (blue), and pixelized points (orange). The evenly distributed case provides the closest approximation to the curve, and remains within Δl of the curve histogram function. Both the randomly generated points and the pixelized randomly generated points stay within $2\Delta l$, and so all three methods work as advertised.

Example 16. Let us also recheck that the global distance histograms are indeed invariant under Euclidean transformations. Since they preserve distances, the net effect of a rigid motion is to resample the curve. In the random and pixelized cases, this effect is eliminated by selecting the points randomly, but in the uniformly distributed case, selecting different points could affect $\lambda_n(r)$.

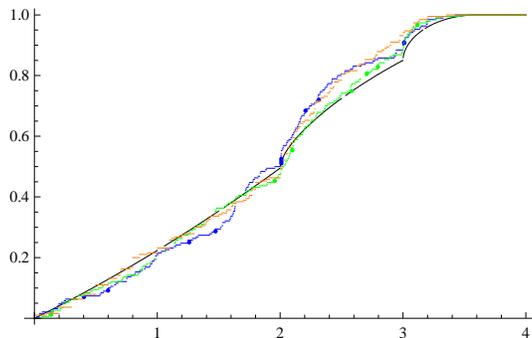


Figure 7: Comparison of Approximate Histograms of a Rectangle.

In Figure 8, we plot the approximate histograms for a square calculated using different offsets for the points. The first graph uses $n = 20$ sample points, while the second has $n = 50$. The individual graphs are clearly different, but as we increase the number of points, they are converging to the same global histogram function (4.8), in accordance with our general results.

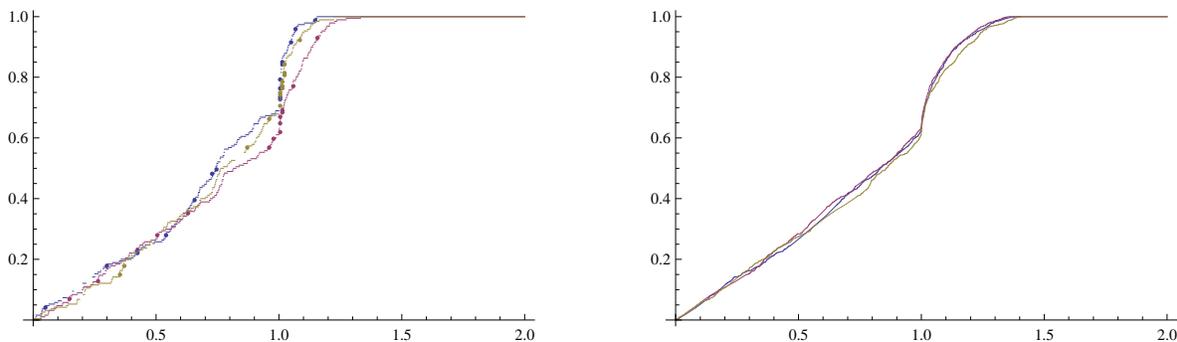


Figure 8: Approximate Histograms of a Square.

The one remaining point is to show how to bound the number of pixels that the curve passes through. surprisingly, we were unable to find a statement of such a result in the literature, and so provide a direct proof. By convention, a pixel is a *closed* square box, and so two neighboring pixels will have a nonempty intersection along a common bounding line segment. As before, a pixel is counted even if the curve only touches one point.

Theorem 17. *Let C be a rectifiable closed curve of length $L = l(C) \geq 3\sqrt{2}w$. Then C passes through at most $n = 3L/(\sqrt{2}w)$ square pixels of pixel width w .*

Proof: By a simple rescaling and translation, we can assume without loss of generality that the pixels have unit width, $w = 1$, with corners on the integer lattice $\mathbb{Z}^2 \subset \mathbb{R}^2$.

We first remark that the bound given in the statement of the theorem is sharp. It is achieved when, for instance, C is a rectangle whose sides are at a 45° angle with the horizontal and whose corners lie on the pixel lattice. For example, the rectangle in Figure 9 has length $L = 14\sqrt{2}$ and goes through $n = 42 = 3L/\sqrt{2}$ pixels. Even if one disallows

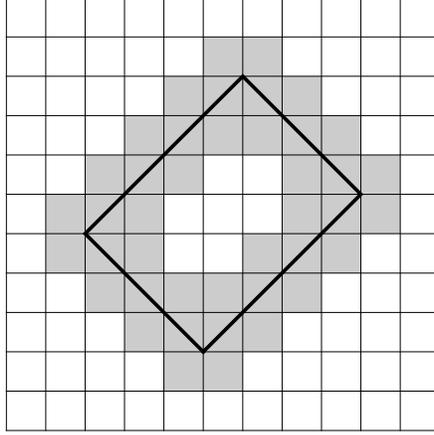


Figure 9: Pixels Containing a Rectangle.

pixels that meet the curve in a single point, one can slightly perturb such rectangles to obtain curves that come arbitrarily close to the pixel bound.

Secondly, if C passes through ≤ 9 pixels, then the bound is trivially satisfied due to the restriction on $l(C)$. Indeed, by our convention, an individual pixel boundary, which forms a unit square of length 4, passes through all $9 > 12/\sqrt{2}$ pixels, and hence the stated bound is not valid for curves of shorter length.

Keep in mind that we are dealing with a closed curve. If C is entirely contained within a *horizontal two pixel strip* $\widehat{H}_j = \{(x, y) \mid j - 1 < y < j + 1\}$ for some $j \in \mathbb{Z}$, then it passes through at most $L + 4$ pixels — the maximum being approached by curves that go arbitrarily closely around the horizontal line segment $\{(x, j) \mid i \leq x \leq i + L/2\}$ for L an even integer. Since $L > 4$, the number of pixels the segment passes through is bounded by $L + 4 < 3L/\sqrt{2}$, proving the theorem in this simple case. Thus, for the remainder of the proof, we may assume C passes through at least 10 pixels and is not contained in a two pixel strip.

Our strategy is to break C up into a finite collection of non-overlapping segments, prove an appropriate bound on the number of pixels each segment passes through, and then successively merge adjacent segments while maintaining the bound. To each such segment $S \subset C$, let $p(S)$ denote the collection of pixels that it passes through. We split $p(S) = q(S) \cup e(S)$ into two disjoint subsets, where $e(S)$ consists of some of the pixels that contain the endpoints of S , satisfying

$$2 \leq \#e(S) \leq 4, \quad \text{while} \quad \#q(S) \leq \frac{3l(S)}{\sqrt{2}}. \quad (6.4)$$

The initial segments will all be of the form $S = E_0 \cup F \cup E_1$, where F is curve segment that, apart from its endpoints, lies strictly within a *horizontal one pixel strip* $H_j = \{(x, y) \mid j < y < j + 1\}$ for some $j \in \mathbb{Z}$, and such that one of its endpoints lies on one bounding horizontal line $y = j$, while the other endpoint lies on the other line $y = j + 1$, while E_0 and E_1 are closed straight line segments contained, respectively, in the horizontal lines $y = j$ and $y = j + 1$, each sharing a common endpoint with F . We allow E_0 and/or E_1 to consist of just a single endpoint of S . We define $e(S)$ to be the set of pixels that contain the endpoints of S and lie outside the strip H_j , while $q(S) = p(S) \setminus e(S)$. Note that each endpoint contributes either two or one pixels to $e(S)$, depending on whether it is a lattice

point or not, and so $2 \leq \#e(S) \leq 4$. A representative example can be seen in Figure 10: F is the curved segment, E_0 is the right hand endpoint, while E_1 is the line segment on the left. The 14 pixels in $q(S)$ are lightly shaded, while the 3 in $e(S)$ are darker.

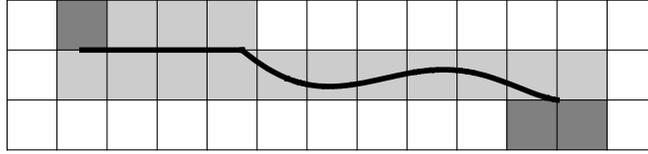


Figure 10: Segment Contained in a Horizontal Strip.

In all situations $l(F) \geq 1$. If F passes through $k \geq 2$ pixels in the strip, then $l(F) \geq \sqrt{1 + (k - 2)^2}$, with the minimal length corresponding to a straight line whose ends are lattice points on opposite sides of the strip. Therefore,

$$k \leq \frac{k l(F)}{\sqrt{1 + (k - 2)^2}} \leq \frac{3 l(F)}{\sqrt{2}},$$

where the maximum is attained when $k = 3$. (Keep in mind that $k \geq 2$ is an integer. In Figure 10, $k = 8$.) The bound trivially holds when $k = 1$. Moreover, it's not hard to see that, by construction, the total number of pixels in $q(C)$ is bounded by

$$\#q(S) \leq k + 2l(E_0) + 2l(E_1) \leq \frac{3}{\sqrt{2}} [l(E_0) + l(E_1) + l(F)] = \frac{3l(S)}{\sqrt{2}}.$$

There are several cases, depending on whether the endpoints of S and F lie on the lattice or not; details are left to the reader.

Now, let's see what happens when we joint two adjacent segments that individually satisfy (6.4), forming a larger segment $S = S_1 \cup S_2 \subsetneq C$. The common endpoint of S_1 and S_2 (which is no longer an endpoint of S) is in either one or two pixels that belong to both $e(S_1)$ and $e(S_2)$, the number depending on whether or not it is a lattice point. If S_1 and S_2 are associated with different horizontal strips, e.g., H_j and H_{j+1} , then these one or two pixels already appear in $q(S) = q(S_1) \cup q(S_2)$, and so the bound (6.4) is an immediate consequence of the bounds for S_1 and S_2 . On the other hand, if S_1 and S_2 are associated with the same horizontal strip, then the one or two pixels in both $e(S_1)$ and $e(S_2)$ that contain the common endpoint do not appear in $q(S_1)$ or $q(S_2)$. However, there are an equal number of pixels that are in both $q(S_1)$ and $q(S_2)$, and so $\#(e(S_1) \cap e(S_2)) = \#(q(S_1) \cap q(S_2))$. Thus, setting $q(S) = q(S_1) \cup q(S_2) \cup (e(S_1) \cap e(S_2))$, we conclude that $\#q(S) \leq \#q(S_1) + \#q(S_2)$, and hence (6.4) also holds for such combined segments.

We continue to merge segments in this fashion. The only detail is in the final merger to form $C = \widehat{S} \cup \widetilde{S}$, when the segments have two common endpoints. If the endpoints are more than a single pixel apart, then the merger process can be completed without interference, and the result follows. (This is where the argument breaks down for a pixel boundary square.) But, our initial assumptions on C guarantee that we can suitably order the merging process to ensure that this is the case. *Q.E.D.*

Remark: For a non-closed curve, the same argument implies the number of pixels is bounded by $n = 4 + 3L/(\sqrt{2}w)$. Again, the bound is tight, as can be seen by deleting a suitable part of the rectangle in Figure 9.

7 Histogram–Based Shape Comparison.

In this section, we discuss the question of whether distance histograms can be used, both practically and theoretically, as a means of distinguishing shapes that are not rigidly equivalent. We begin with the practical aspects. As we know, if two curves have different global histogram functions, they cannot be rigidly equivalent. For curves arising from digital images, we will approximate the global histogram function by its discrete approximation based on a reasonably dense sampling of the curve. Since the precision of the approximations is proportional to $\Delta l = L/n$, we will calculate the average difference between two histogram plots, normalized with respect to Δl . Our assumption is that differences less than 1 represent histogram approximations that cannot be distinguished.

Two tables of these values for a few elementary sample shapes are displayed below. We use random point distributions to illustrate that identical parameterizations do not necessarily give identical sample histograms. This is also evident from the fact that the matrix is not symmetric — different random points were chosen for each trial. However, symmetrically placed entries generally correlate highly, indicating that the comparison is working as intended.

In the first table, we discretize using only $n = 20$ points. In this case, there is too small a sample set to be able to clearly distinguish the shapes. Indeed, the 2×3 rectangle and the star appear more similar to each other than they are to a second randomized version of themselves. However, if we are trying to determine whether or not the parameterization for the star and the circle were the same, a value of $5.39 \Delta l$ is reasonably strong evidence that they are different.

Shape	(a)	(b)	(c)	(d)	(e)	(f)
(a) triangle	.35	1.16	1.46	4.20	2.36	3.16
(b) square	1.45	.51	3.63	2.46	1.59	2.89
(c) circle	3.65	4.17	.67	5.87	3.14	5.39
(d) 2×3 rect.	3.85	1.95	4.82	1.78	1.85	.72
(e) 1×3 rect.	1.10	1.86	4.02	2.31	1.25	1.93
(f) star	3.90	3.80	5.75	.72	2.55	1.22

20 Point Comparison Matrix

As we increase the number of points and the precision, the computation time increases (in proportion to n^2 for calculating the histograms and n for comparing them), but the ability to differentiate shapes increases as well. In the corresponding table for $n = 500$ points, it is now clear that none of the shapes are rigidly equivalent to any of the others. The value of 4 for comparing the 1×3 rectangle to itself is slightly high, but it is still significantly less

than any of the values for comparing two different shapes.

Shape	(a)	(b)	(c)	(d)	(e)	(f)
(a) triangle	2.25	20.4	66.9	81.0	28.5	76.8
(b) square	28.2	.54	81.2	73.6	34.8	72.1
(c) circle	66.9	79.6	.53	137.	89.2	138.
(d) 2×3 rect.	85.8	75.9	141.	2.15	53.4	9.86
(e) 1×3 rect.	31.8	36.7	83.7	55.7	4.00	46.5
(f) star	81.0	74.3	139.	9.31	60.5	.94

500 Point Comparison Matrix

The ability to distinguish our sample shapes begs the question as to whether or not all shapes can be distinguished by their histograms. As we stated earlier, while almost all finite sets of points can be reconstructed from the distances between them, there are counterexamples, including the kite and trapezoid shown in Figure 1. The distance histogram based on the corner points will not distinguish between them. However, the curve histograms $H_C(r)$ based on the two outer polygons can easily be distinguished. In Figure 11 we plot the approximate global histograms $\Lambda_n(r)$ based on $n = 20$ points. The kite is shown in blue and the trapezoid is shown in purple.

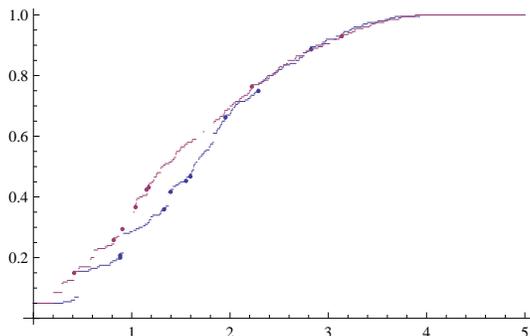


Figure 11: Curve Histograms for the Kite and Trapezoid.

The key theoretical issue, then, is whether distance histograms uniquely prescribe shapes up to rigid motion. While Boutin and Kemper’s result [5, 6] shows that this is not the case for finite point sets, there are indications that it is true for curve histograms, and so we state the desired result as a conjecture.

Conjecture: Let C and \tilde{C} be two fully regular plane curves. Then $H_C(r) = H_{\tilde{C}}(r)$ for all $r \geq 0$ if and only if the curves are rigidly equivalent: $C \simeq \tilde{C}$.

While we as yet have been unable to construct a complete proof, let us present some additional evidence in its favor. One evident proof strategy would be to approximate the histograms by sampling and apply our Convergence Theorem 4. If one could show that the sample points do not, at least when taken sufficiently densely along the curve, lie in the exceptional set of Theorem 2, then our conjecture would follow. A second strategy would be

to first prove the result for polygons, based on our observations that, even when their corners lie in the exceptional set, their curve histograms still appear to uniquely characterize them. Again, the Convergence Theorem would then establish the conjecture.

Barring a resolution of the latter problem, let us discuss what properties of the curve C can be gleaned from its histogram. First of all, the curve's diameter is equal to the minimal value of r for which $H_C(r) = 1$. Secondly, values where the derivative of the histogram function is very large usually have geometric significance. In the square histogram — see Figures 6 and 8 — this occurs at $r = 1$. In polygons, such values often correspond to distances between parallel sides, because, at such a distance, the disk centered on one of the parallel sides suddenly begins to contain points on the opposite side. For shapes with multiple pairs of parallel sides, we can see this effect at several values of r — such as when $r = 2$ and $r = 3$ in the case of a 2×3 rectangle shown in Figure 7. The effect becomes more visible as n increases since the amount of noise decreases. The size of the effect depends on the overall length of the parallel sides; for instance, the slope at $r = 3$ in Figure 7 is larger than that at $r = 2$. However, not every value where the derivative is large is the result of such parallel sides — the kite from the pathological Boutin–Kemper example shown in Figure 11 has two visible jumps, but no parallel sides.

In a more theoretical direction, let us compute the Taylor expansion of the histogram function $H_C(r)$ at $r = 0$, assuming that C is sufficiently smooth. We begin by expanding the local histogram function based at a point $z \in C$. To expedite the analysis, we apply a suitable rigid motion to move the curve into a “normal form” so that z is at the origin, and the tangent at z is horizontal. Thus, in a neighborhood of $z = (0, 0)$, the curve is the graph of a function $y = y(x)$ with $y(0) = 0$ and $y'(0) = 0$. As a consequence of the moving frame recurrence formulae developed in [9] — or working by direct analysis — we can write down the following Taylor expansion.

Lemma 18. *Under the above assumptions,*

$$y = \frac{1}{2} \kappa x^2 + \frac{1}{6} \kappa_s x^3 + \frac{1}{24} (\kappa_{ss} + 3\kappa^3) x^4 + \frac{1}{120} (\kappa_{sss} + 19\kappa^2 \kappa_s) x^5 + \dots, \quad (7.1)$$

where $\kappa, \kappa_s, \kappa_{ss}, \dots$ denote, respectively, the curvature and its successive arc length derivatives evaluated at $z = (0, 0)$.

We use this formula to find a Taylor expansion for the local histogram function $h_C(r, z)$ at $r = 0$. Assume that $r \ll 1$ is small. The curve (7.1) will intersect the circle of radius r centered at the origin at two points $z_{\pm} = (x_{\pm}, y_{\pm}) = (x_{\pm}, y(x_{\pm}))$, which are the two solutions to the quadratic equation

$$x^2 + y(x)^2 = r^2.$$

Substituting the expansion (7.1) and solving the resulting series equation for x , we find

$$\begin{aligned} x_+ &= r - \frac{1}{8} \kappa^2 r^3 - \frac{1}{12} \kappa \kappa_s r^4 - \left(\frac{1}{48} \kappa \kappa_{ss} + \frac{1}{72} \kappa_s^2 + \frac{1}{128} \kappa^4 \right) r^5 + \dots, \\ x_- &= -r + \frac{1}{8} \kappa^2 r^3 - \frac{1}{12} \kappa \kappa_s r^4 + \left(\frac{1}{48} \kappa \kappa_{ss} + \frac{1}{72} \kappa_s^2 + \frac{1}{128} \kappa^4 \right) r^5 + \dots. \end{aligned} \quad (7.2)$$

Thus, again using (7.1),

$$\begin{aligned}
h_C(r, z) &= \int_{x_+}^{x_-} \sqrt{1 + y'(x)^2} dx \\
&= \int_{x_+}^{x_-} \sqrt{1 + \kappa^2 x^2 + \kappa \kappa_s x^3 + \left(\frac{1}{3} \kappa \kappa_{ss} + \frac{1}{4} \kappa_s^2 + \kappa^4\right) x^4 + \dots} dx \\
&= \int_{x_+}^{x_-} \left[1 + \frac{1}{2} \kappa^2 x^2 + \frac{1}{2} \kappa \kappa_s x^3 + \left(\frac{1}{6} \kappa \kappa_{ss} + \frac{1}{8} \kappa_s^2 + \frac{3}{8} \kappa^4\right) x^4 + \dots \right] dx \\
&= \left[x_+ + \frac{1}{6} \kappa^2 x_+^3 + \frac{1}{8} \kappa \kappa_s x_+^4 + \left(\frac{1}{30} \kappa \kappa_{ss} + \frac{1}{40} \kappa_s^2 + \frac{3}{40} \kappa^4\right) x_+^5 + \dots \right] - \\
&\quad - \left[x_- + \frac{1}{6} \kappa^2 x_-^3 + \frac{1}{8} \kappa \kappa_s x_-^4 + \left(\frac{1}{30} \kappa \kappa_{ss} + \frac{1}{40} \kappa_s^2 + \frac{3}{40} \kappa^4\right) x_-^5 + \dots \right].
\end{aligned}$$

We now substitute (7.2) to deduce that

$$\begin{aligned}
h_C(r, z) &= \left(r + \frac{1}{24} \kappa^2 r^3 + \frac{1}{24} \kappa \kappa_s r^4 + \left(\frac{1}{80} \kappa \kappa_{ss} + \frac{1}{90} \kappa_s^2 + \frac{3}{640} \kappa^4\right) r^5 + \dots \right) - \\
&\quad - \left(-r - \frac{1}{24} \kappa^2 r^3 + \frac{1}{24} \kappa \kappa_s r^4 - \left(\frac{1}{80} \kappa \kappa_{ss} + \frac{1}{90} \kappa_s^2 + \frac{3}{640} \kappa^4\right) r^5 + \dots \right) \quad (7.3) \\
&= 2r + \frac{1}{12} \kappa^2 r^3 + \left(\frac{1}{40} \kappa \kappa_{ss} + \frac{1}{45} \kappa_s^2 + \frac{3}{320} \kappa^4\right) r^5 + \dots .
\end{aligned}$$

Invariance of both sides of this formula under rigid motions implies that it holds as written at any point $z \in C$.

To obtain the Taylor expansion of the global histogram function, we substitute (7.3) back into (2.9), resulting in

$$H_C(r) = 2r + \frac{r^3}{12L} \int_C \kappa^2 ds + \frac{r^5}{5L} \int_C \left(\frac{1}{8} \kappa \kappa_{ss} + \frac{1}{9} \kappa_s^2 + \frac{3}{64} \kappa^4\right) ds + \dots \quad (7.4)$$

If C is a closed curve, then we can integrate by parts to simplify the final integral:

$$H_C(r) = 2r + \frac{r^3}{12L} \oint_C \kappa^2 ds + \frac{r^5}{40L} \oint_C \left(\frac{3}{8} \kappa^4 - \frac{1}{9} \kappa_s^2\right) ds + \dots \quad (7.5)$$

Each integral appearing in the Taylor expansion is uniquely determined by the histogram function. An interesting question is whether the resulting integral moments, depending on curvature and its arc length derivatives, uniquely prescribe the curve up to rigid motion. If so, this would prove our Conjecture for regular curves.

As noted above, another direction worth investigating further is the distance histogram function of a polygon. Indeed, if one can prove that the global distance histogram of a simple closed polygon (as opposed to the discrete histogram based on its corners) uniquely characterizes the polygon up to rigid motion, then our conjecture for general curves would follow by suitably approximating them by their interpolating polygons. Some evidence in this direction is as follows.

Let K be a simple closed polygon. Let $l_\star > 0$ be the minimum side length, and $d_\star > 0$ be the minimum distance between any two non-adjacent sides. Set $m_\star = \min\{l_\star, d_\star\} > 0$. Then any disk $B_r(z)$ centered at a point $z \in K$ of radius $0 < r < \frac{1}{2} m_\star$ intersects K in either one or two sides, the latter possibility only occurring when z is within a distance r of the nearest corner. Let z_1, \dots, z_n be the corners of K and let θ_j denote the interior angle at z_j — see Figure 12. Then,

$$h_K(r, z) = l(K \cap B_r(z)) = \begin{cases} x_j + y_j + r, & x_j = d(z, z_j) < r, \\ 2r, & \text{otherwise,} \end{cases} \quad z \in K, \quad r < \frac{1}{2} m_\star, \quad (7.6)$$

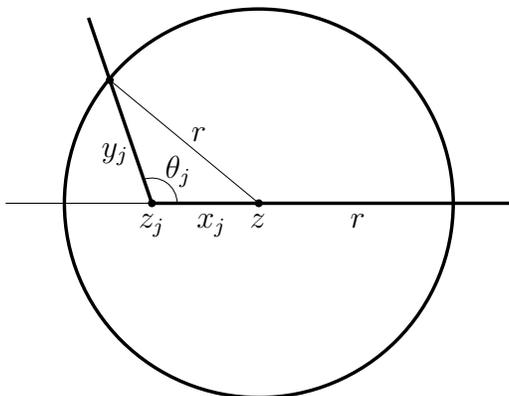


Figure 12: Intersection of a Polygon and a Disk.

where, by the Law of Cosines, y_j solves the quadratic equation

$$y_j^2 - 2x_j y_j \cos \theta_j + x_j^2 = r^2, \quad \text{with} \quad x_j = d(z, z_j) < r. \quad (7.7)$$

Thus, when $r < \frac{1}{2}m_*$, the global histogram function (2.9) for the polygon takes the form

$$H_K(r) = \frac{1}{L} \int_K h_K(r, z(s)) ds = 2r - \frac{2nr^2}{L} + \frac{2}{L} \sum_{j=1}^n \Psi(\theta_j, r), \quad (7.8)$$

where

$$\Psi(\theta_j, r) = \int_0^r [x + y_j(x)] dx, \quad (7.9)$$

with $y_j = y_j(x)$ for $x = x_j$ implicitly defined by (7.7). (There is, in fact, an explicit, but not very enlightening, formula for this integral in terms of elementary functions.)

Observe that (7.8) is a *symmetric function* of the polygonal angles $\theta_1, \dots, \theta_n$, i.e., it is not affected by permutations thereof. Moreover, for distinct angles $\theta_j \neq \theta_k$, the integrals $\Psi(\theta_j, r)$ are linearly independent functions of r . This implies that one can recover the set of polygonal angles $\{\theta_1, \dots, \theta_n\}$ from knowledge of the global histogram function $H_K(r)$ for $0 < r < \frac{1}{2}m_*$. In other words, the polygon's global histogram function does determine its angles up to a permutation.

The strategy for continuing a possible proof would be to gradually increase the size of r . Since, for small r , the histogram function has determined the angles, its form is fixed for all $r \leq \frac{1}{2}m_*$. For $r > \frac{1}{2}m_*$, the functional form will change, and this will serve to prescribe m_* , the minimal side length or minimal distance between non-adjacent sides. Proceeding in this fashion, as r gradually increases, more and more sides of the polygon can be covered by a disk of that radius, providing more and more geometric information about the polygon from the resulting histogram. This points the way to a proof of the polygonal histogram conjecture, and hence the curve conjecture. However, rigorously implementing such a proof strategy appears to be quite intricate, and the details remain to be sorted out.

8 Extensions.

There are a number of directions in which this research could be extended. The most obvious is to apply it to more substantial practical problems to gauge whether or not it can compete with other methods of object recognition, particularly its ability to handle noisy images. Other potential avenues for future research involve extending the ideas and techniques to recognize shapes under more general group actions, and to apply these methods to curves, surfaces and other subsets of higher dimensional spaces.

8.1 Higher Dimensions

We can easily extend our analysis to objects in three or more dimensions with minimal changes in the methodology. Local and global histograms of curves in \mathbb{R}^3 are defined by simply replacing the disk of radius r by the solid ball of that radius in the formulas (2.8), (2.9). For example, consider the saddle-like curve parametrized by

$$z(t) = (\cos t, \sin t, \cos 2t), \quad 0 \leq t \leq 2\pi. \quad (8.1)$$

In Figure 13, we plot the discrete approximations $\Lambda_n(r)$ to the curve histogram. The blue plot corresponds to $n = 10$ points, purple to $n = 20$, and yellow $n = 30$. Note that the discrete cumulative histograms appear to converge as $n \rightarrow \infty$.

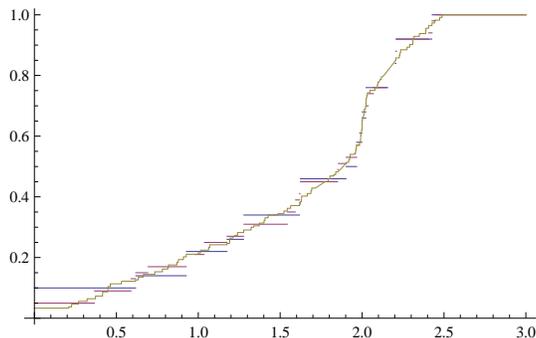


Figure 13: Approximate Distance Histograms for the Three-Dimensional Saddle Curve

The next extension is to apply the histogram method to two-dimensional surfaces in three-dimensional space. We consider the case of piecewise smooth surfaces $S \subset \mathbb{R}^3$ with finite area $A = \text{area}(S) < \infty$. Let $P_n \subset S$ be a finite set of n sample points. We retain the meaning of $\lambda_n(r, z)$ as the proportion of points within a distance r of the point z , (2.4), and $\Lambda_n(r)$ as its average, (2.6).

We claim without proof that these approximate the corresponding local and global surface histogram functions

$$h_S(r, z) = \text{Area}(S \cap B_r(z)), \quad H_S(r) = \frac{1}{A} \iint_S h_S(r, z) dS. \quad (8.2)$$

The convergence of the discrete histograms is illustrated in Figure 14. Plots of the discrete approximations $\Lambda_n(r)$ for the unit sphere $S^2 = \{\|z\| = 1\} \subset \mathbb{R}^3$ are shown, with $n = 10$ in

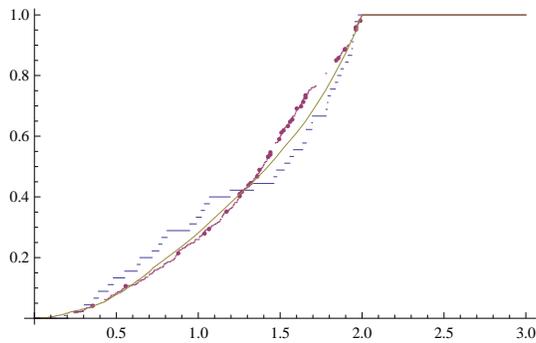


Figure 14: Approximate Distance Histograms of a Sphere.

blue, $n = 30$ in purple, and $n = 100$ in yellow. The approximations appear to be converging as $n \rightarrow \infty$, albeit at a slower rate than was the case with curves.

Potential future work could involve prove a convergence theorem for surfaces and higher dimensional submanifolds of Euclidean space, along the lines of Theorem 4. A more challenging question is whether distance histograms can be used to distinguish subsets of differing dimensions. Or, to state this another way, can one determine the dimension of a subset from some innate property of its histogram?

8.2 Area Histograms

In computer vision applications, the invariance of objects under the *equi-affine group*, consisting of all area- or volume-preserving affine transformations of \mathbb{R}^n , is of great importance; see, for instance, [7, 11, 18]. Equi-affine transformations can be viewed as approximations to projective transformations, valid for moderately tilted planar objects, e.g., a plate lying on a table when viewed at an angle. The basic planar equi-affine joint invariant is the area of a triangle formed by three points, and hence the histogram formed by the areas of the triangles of all triples in a finite point configuration is invariant under the equi-affine group. Boutin and Kemper, [5], also proved that generic planar point configurations are uniquely determined, up to equi-affine transformations, by their area histograms, but, just as with distance histograms, there is a lower dimensional algebraic subvariety of exceptional configurations. For us, the key question is convergence of the cumulative area histogram based on sample points on a plane curve.

To define the continuous area histogram for a curve, we first note that the distance histogram function (2.9) can be expressed in the alternative form

$$H_C(r) = \frac{1}{L} \int_C \int_C \sigma_r(d(z(s), z(s'))) ds ds', \quad (8.3)$$

where

$$\sigma_r(x) = \begin{cases} 1, & x < r, \\ 0, & x \geq r \end{cases} \quad (8.4)$$

is the unit step function based at $x = r$. By analogy, we define the area histogram function

$$A_C(r) = \frac{1}{L^3} \int_C \int_C \int_C \sigma_r(\text{Area}(z(s), z(s'), z(s''))) ds ds' ds'', \quad (8.5)$$

where s, s', s'' all refer to the *equi-affine arc length* of the curve, [10], while $L = \int_C ds$ is its total equi-affine arc length. In local coordinates, if the curve is the graph of a function $y(x)$ then the equi-affine arc length element is given by $ds = \sqrt[3]{y''(x)} dx$.

We claim, without proof, that the analogous approximate cumulative area histogram is

$$A_P(r) = \frac{1}{n(n-1)(n-2)} \sum_{z \neq z' \neq z'' \in P} \sigma_r(\text{Area}(z, z', z'')). \quad (8.6)$$

We calculate and plot this quantity, and observe that it does converge to the area histogram function (8.5) as $n \rightarrow \infty$. Figure 15 illustrates the convergence for a circle, taking $n = 10$ points in blue, $n = 20$ in purple, and $n = 30$ in yellow.

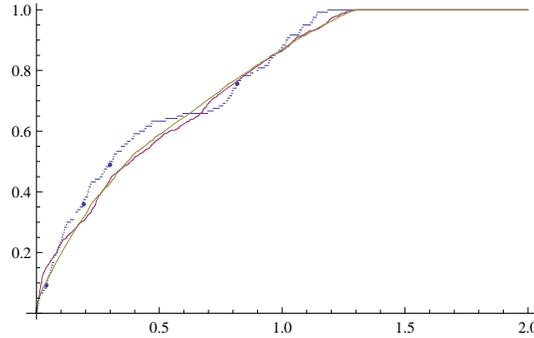


Figure 15: Area Histogram of a Circle.

Let us end with an investigation of the equi-affine invariance of the curve histogram function. Since rectangles of the same area are equivalent under a equi-affine transformation, they have identical area histograms. In Figure 16, we plot area histograms for a 2×2 square in blue, a 1×4 rectangle in purple, and a $.5 \times 8$ rectangle in yellow. Indeed, with $n = 30$ sample points, the graphs are quite close.

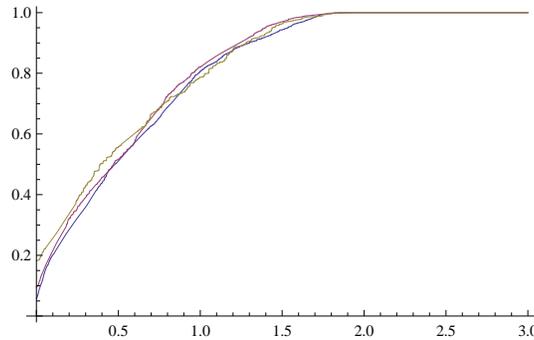


Figure 16: Area Histograms of Affine-Equivalent Rectangles

We defer further development of these ideas to future projects.

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